

# Supplemental Notes

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3 types of proof

1. Direct :  $P \rightarrow Q$

assume  $P$ , derive  $Q$

2. Indirect : "reductio ad absurdum" (reduce to absurdity)

assume  $P \wedge \sim Q$ , derive  $C \wedge \sim C$

$$P \rightarrow Q \stackrel{\text{M.I.}}{=} \sim P \vee Q = \sim(P \wedge \sim Q)$$

$$\therefore \sim(P \rightarrow Q) = \boxed{P \wedge \sim Q}$$

Thm:  $(P \rightarrow (Q \wedge \sim Q)) \rightarrow \sim P$ .

P	Q	$P \rightarrow (Q \wedge \sim Q)$			$\Rightarrow$	$\sim P$
1	1	1	0	0	1	0
0	1	0	1	0	1	1
1	0	1	0	0	1	0
0	0	0	1	0	1	1

∴ Valid

Ex. Thm:  $\forall \epsilon > 0 \quad a \leq b + \epsilon \rightarrow a \leq b.$

Pf. Suppose not. i.e.  $\forall \epsilon > 0: (a \leq b + \epsilon) \wedge (a > b)$

$$a = \frac{a}{2} + \frac{a}{2} \stackrel{\text{Assum.}}{>} \frac{a}{2} + \frac{b}{2} = b - \frac{b}{2} + \frac{a}{2} = b + \left\{ \frac{a-b}{2} \right\} \epsilon' > 0 \quad \text{since:}$$
$$a > b \iff a - b > 0$$
$$= b + \epsilon' \wedge \epsilon' > 0 \quad \therefore \frac{a-b}{2} > 0$$
$$\therefore \epsilon' > 0$$

$$\therefore a > b + \epsilon' \quad \text{w/ } \epsilon' > 0$$

 contradiction: since  $a \leq b + \epsilon \quad \forall \epsilon > 0$

Thm.  $Q \wedge \sim Q \rightarrow P$

(true AND false imply anything)

# IMAC example: finite sigma-algebra

Note: HW1/2

Sample space  $X$  contains four elements  $X = \{a, b, c, d\}$ .  
So its power set or set of subsets  $2^X$  contains  $2^4$  elements. Thus the power set  $2^X$  gives rise to  $2^{16}$  subcollections of sets. Suppose that  $\mathcal{A} \subset 2^X$  is one of these  $2^{16}$  set subcollections of  $2^X$  and that  $\mathcal{A}$  contains the following 7 sets:

$$\mathcal{A} = \{ \emptyset, \{a\}, \{b\}, \{a, b\}, \{a, c, d\}, \{b, c, d\}, X \}.$$

Can we define a probability space  $(X, \mathcal{A}, P)$  using the set collection  $\mathcal{A}$  if  $P$  is a probability measure?

(use IMAC format to answer)

Soln:

I: sigma-algebra (s.a.)

M:  $\mathcal{Q} = 2^X$  is a sigma-algebra

iff  $\mathcal{Q}$  is CUT:

$$\left\{ \begin{array}{l} C: A \in \mathcal{Q} \implies A^c \in \mathcal{Q} \\ U: A_1 \in \mathcal{Q}, A_2 \in \mathcal{Q}, \dots \implies \bigcup_{k=1}^{\infty} A_k \in \mathcal{Q} \\ T: X \in \mathcal{Q} \end{array} \right.$$

A:  $\{a\}^c = \{b, c, d\} \in \mathcal{Q}$

C:  $\{b\}^c = \{a, c, d\} \in \mathcal{Q}$ .

$\{a, b\}^c = \{c, d\} \notin \mathcal{Q}$

$$\left. \begin{array}{l} \{b, c, d\}^c = \{a\} \in \mathcal{Q} \\ \{a, c, d\}^c = \{b\} \in \mathcal{Q} \\ \emptyset^c = X \in \mathcal{Q}, X^c = \emptyset \in \mathcal{Q} \end{array} \right\}$$

$\therefore$   $\mathcal{Q}$  is not closed under complement

because  $\{a, b\}^c \notin \mathcal{Q}$

u:  $\{a\} \cup \{b\} = \{a, b\} \in \mathcal{Q}$

$$\{a\} \cup \{a, b\} = \{a, b\} \in \mathcal{Q}$$

$$\{a\} \cup \{a, c, d\} = \{a, c, d\} \in \mathcal{Q}$$

$$\{a\} \cup \{b, c, d\} = X \in \mathcal{Q}$$

$$\{b\} \cup \{a, b\} = \{a, b\} \in \mathcal{Q}$$

$$\{b\} \cup \{a, c, d\} = X \in \mathcal{Q}$$

$$\{b\} \cup \{b, c, d\} = \{b, c, d\} \in \mathcal{Q}$$

$$\emptyset \cup S = S \in \mathcal{Q} \quad \text{if } S \in \mathcal{Q}$$

$$X \cup S = X \in \mathcal{Q} \quad \text{if } S \in \mathcal{Q}$$

$\therefore$   $\mathcal{Q}$  is closed under union.

T:  $X \in \mathcal{Q}$ .

$\therefore$   $\mathcal{Q}$  is ~~not~~  $\sigma$ -algebra

$\therefore$   $\mathcal{Q}$  is not a sigma-algebra.

C: **NO.**

$(X, \mathcal{A}, P)$  cannot be a probability space  
because  $\mathcal{A}$  is not a sigma-algebra.

## Discrete Sample Spaces

Finite or denumerable outcomes

↳ can put into 1-to-1 correspondance w/  $\mathbb{N}$

Ex:  $\Omega = \{a_1, a_2, \dots, a_n\}$  finite sample space

Consider "equally likely outcomes"

$$P[\{a_i\}] = \frac{1}{n} \quad \forall i=1, \dots, n$$

If event consists of  $k$  distinct outcomes:

$$\begin{aligned} P[B] &= P[\{a'_1, a'_2, \dots, a'_k\}] = P[\{a'_1\}] + \dots + P[\{a'_k\}] \\ &= \frac{k}{n} \end{aligned}$$

Ex: 3 coin flips  $\therefore$  8 outcomes

$$\Omega = \{HHH, HHT, HTH, THH, HTT, THT, TTH, TTT\}$$

Assume equiprobable:  $P[xxx] = \frac{1}{8}$

$$P[2 \text{ heads}] = P[HHT \text{ or } HTH \text{ or } THH] = \frac{3}{8}$$

## Counting methods and Combinatorics

In finite sample space w/ equiprobable outcomes.

$$P[\text{event}] = \frac{\text{distinct \# of outcomes that match event}}{\text{total \# distinct outcomes}}$$

Ex: multiple choice exam, 10 questions, 4 options each

$$\# \text{ outcomes} = \underbrace{4 \times 4 \times \dots \times 4}_{10} = 2^{20} \approx 1 \text{ million}$$

## Computer Problem: Birthday Paradox

Q: How large of group of  $k$  people to have prob. at least 0.5 of having at least two people with same birthday

(assume: 365 days/year, independent birthdays, equally likely per day)

Let  $A_k$  = event at least 2 people in group w/  $k$  share a birthday.

define  $q_k = P[A_k^c]$   
↳ no common birthday in group of  $k$ .

$$q(1) = 1$$

$$q(2) = q(1) \cdot \frac{364}{365} = \frac{365}{365} \cdot \frac{364}{365} = \frac{364}{365}$$

$$q(3) = q(2) \cdot \frac{363}{365} = \frac{365}{365} \cdot \frac{364}{365} \cdot \frac{363}{365}$$

⋮

$$q(k) = q(k-1) \cdot \frac{365-k+1}{365} = \frac{365 \cdot 364 \cdot \dots \cdot (365-k+1)}{365^k} = \frac{(365)_k}{365^k}$$

"365 fall  $k$ "  
↓

Computer:  $q(22) = 0.524$

$$q(23) = 0.493$$

∴  $k \geq 23$  people.

## Random sampling (ordered)

- with replacement

- choose  $k$  from  $n$  distinct elements.
- select 1 at a time and "replace" after each selection

$\therefore$  duplicate samples possible

# outcomes: ordered  $k$ -tuple w/  $n$ -options for each position

$$= \underbrace{n \cdot n \cdots n}_k = n^k$$

Ex: multiple choice exam w/ 4-options per question.

- without replacement

- choose  $k$  from  $n$  distinct elements
- select 1 at a time a "remove" after each selecting

$\therefore$  no duplicate samples possible

$$n_1 = n$$

$$n_2 = n-1$$

$$n_3 = n-2$$

$\vdots$

$$n_k = n-k+1$$

$$\therefore \# \text{ outcomes} = n \cdot (n-1)(n-2) \cdots (n-k+1)$$

(note  $k \leq n$ )

Ex: Drawing balls # 1-60 from urn, don't replace

Special case: permutation.

$\hookrightarrow$  Sampling w/o replacement,  $k=n$

$$\# \text{ outcomes} = n \cdot (n-1)(n-2) \cdots (2)(1)$$

$$= \boxed{n!}$$

Defn: Permutation = Self-Bijection

$$- f: S \rightarrow S$$

$$- f: 1\text{-to-1} \quad (\text{injection})$$

$$- f: \text{onto} \quad (\text{surjection})$$

Fundamental Rule of counting

$$\text{arrangement} = \boxed{n_1 \mid n_2 \mid \dots \mid n_k}$$

$$\therefore |\text{arrangements}| = \prod_{k=1}^n n_k$$

$$(\text{=} n^k \text{ if } n_k = n \ \forall k)$$

$P(n, k) = \#$  permutations of length  $k$  on  $n$  elements

$$n \cdot (n-1) \cdots (n - (k+1)) \cdot \frac{(n-k)!}{(n-k)!}$$

$$= \frac{n!}{(n-k)!} \quad \text{for } k \leq n.$$

Defn:  $n! = n(n-1) \cdots 3 \cdot 2 \cdot 1$  for  $n \in \{0, 1, 2, \dots\}$ .

$$(0! = 1)$$

Fact: (Stirling's Approximation)

$$\ln n! \approx n \cdot \ln n - n$$

$$\approx n \cdot \ln n$$

Ex:  $n! < n^n$ , for  $n > 2$

$$\therefore \ln n! < \ln n^n = n \cdot \ln n$$

$$e^n = \sum_{k=0}^{\infty} \frac{n^k}{k!} > \frac{n^n}{n!} \quad \leftarrow \text{since } \sum_{k=0}^{\infty} \frac{n^k}{k!} = \frac{n^0}{0!} + \frac{n^1}{1!} + \frac{n^2}{2!} + \dots + \frac{n^n}{n!} + \dots$$

$$\therefore n! > \frac{n^n}{e^n}$$

$$\therefore \ln n! > n \cdot \ln n - n$$

$$\therefore n \cdot \ln n - n < \ln n! < n \cdot \ln n$$

$$\therefore 1 - \frac{1}{\ln n} < \frac{\ln n!}{n \cdot \ln n} < 1$$

$$\therefore \lim_{n \rightarrow \infty} \frac{\ln n!}{n \cdot \ln n} = 1$$

$$\therefore \ln n! \sim n \cdot \ln n$$

Note:  $\ln n! = \ln \prod_{k=1}^n k = \sum_{k=1}^n \ln k \sim \int_{x=1}^n \ln x \, dx$

pick  $u = \ln x$        $dv = dx$

$du = \frac{1}{x}$        $v = x$

$$= x \cdot \ln x \Big|_{x=1}^{x=n} - x \Big|_{x=1}^{x=n} = n \cdot \ln n - n + 1 \sim n \cdot \ln n - n$$

# Unordered Sampling

order doesn't matter, i.e.,  $(a, b) \sim (b, a)$

- without replacement

- choose  $k$  from  $n$  w/o replacement
- outcome order does not matter

Ex: "Standard lottery"

Draw  $k$  balls "lottery numbers"

any order wins

$\therefore$  60 balls choose  $k$ :  $60 \cdot 59 \dots 55 = \frac{60!}{54!}$

*(Annotations:  $n!$  points to 60!,  $(n-k)!$  points to 54!)*

but each outcome has:  $k \cdot 5 \cdot 4 \cdot 3 \cdot 2 \cdot 1 = k!$  equivalent outcomes

$\therefore$  # outcomes =  $\frac{60!}{54! k!}$

Defn:  $\frac{n!}{(n-k)!k!} = \binom{n}{k}$  is the binomial coefficient

- with replacement

- choose  $k$  from  $n$  w/ replacement.
- outcome order does not matter.

Ex: "Pizza toppings"

$k \rightarrow$  4 toppings  
 $n \rightarrow$  5 choices  
 (duplicate OK)

A	B	C	D
XX		X	XX
X	XX	X	X

$\leftarrow$  outcome #1  
 $\leftarrow$  outcome #2

Shorthand:

XX || X | XX #1  
 X | XX | X | X #2

5 "X" + 3 " |"  
 "stars and bars"

$\therefore$  every outcome unordered w/ replacement is equivalent to a unique ordered sample of " $X$ " and " $1$ "

$\therefore$  # distinct outcomes: sample w/o replacement from urn w/  $k$ -black balls and  $(n-1)$  white balls.  
 ( $\therefore$  total =  $n-1+k$  balls)

Ex: (cont) How many distinct permutations of  $k$  black balls and  $n-1$  white balls? ordered

Consider balls #  $1, 2, \dots, n-1+k$  (no color)  
 then choose  $k$ , put " $X$ " there. " $1$ " else.

$$\frac{(n-1+k)!}{k! (n-1)!} \binom{n-1+k}{k} = \binom{n-1+k}{n-1} \frac{(n-1+k)!}{(n-1)! k!}$$

$\uparrow$  choose  $k$  for " $X$ " pos.
 $\uparrow$  choose  $n-1$  for " $1$ " pos.

### Counting method summary

		Replacement	
		with	without
Ordered	yes	$n^k$ <u>Ex:</u> # ways to answer multiple choice test	$n(n-1)\dots(n-(k-1))$ <u>Ex:</u> How many ways to rank $k$ students in class with $n$ -students
	no	$\frac{(n-1+k)!}{(n-1)! k!}$ <u>Ex:</u> # different pizzas w/ $k$ -toppings (duplicate OK)	$\frac{n!}{k!(n-k)!}$ <u>Ex:</u> # ways to draw " $k$ " lottery numbers

$\rightarrow \binom{n}{k}$

$$\binom{n-1+k}{n-1} = \binom{n-1+k}{k}$$

# Binomial Coefficients

$P(n, k) =$  # permutations of length  $k$  on  $n$  elements

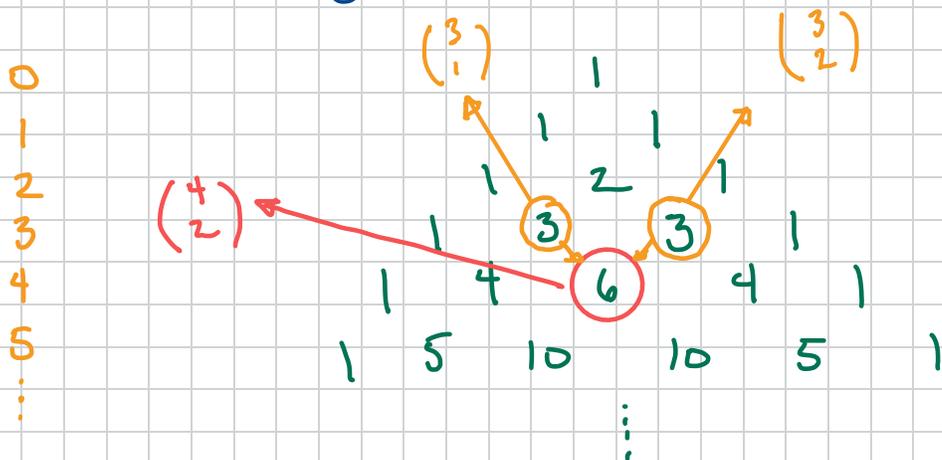
$$= \frac{n!}{(n-k)!} \quad \text{for } k \leq n.$$

$C(n, k) =$  # combinations of length  $k$  on  $n$  elements

$$= \binom{n}{k} = \frac{n!}{(n-k)! k!} \quad \text{for } k \leq n.$$

Note:  $\binom{n}{0} = 1$   $\longleftrightarrow \emptyset$   
 $\binom{n}{n} = 1$   $\longleftrightarrow \Omega$

## Pascal's triangle



$$\begin{aligned} (x+y)^0 &= 1 \\ (x+y)^1 &= x+y \\ (x+y)^2 &= x^2+2xy+y^2 \\ (x+y)^3 &= x^3+3x^2y+3xy^2+y^3 \\ (x+y)^4 &= \dots \\ (x+y)^5 &= \dots \end{aligned}$$

note: # ways of choosing

w/  $2^x$  and  $1^y$

just a hypothesis now  
proof is binomial theorem

$$\text{for } \binom{n}{k} = \frac{n!}{k!(n-k)!} = \frac{n!}{(n-k)! k!} = \binom{n}{n-k}$$

Thm: (Pascal's Formula)

$$\binom{n}{k} = \binom{n-1}{k-1} + \binom{n-1}{k}$$

Prf:  $\binom{n-1}{k-1} + \binom{n-1}{k} = \frac{(n-1)!}{(k-1)! (n-1-(k-1))!} + \frac{(n-1)!}{k! (n-1-k)!}$

$$= \frac{k}{k} \cdot \frac{(n-1)!}{(k-1)! (n-k)!} + \frac{(n-1)!}{k! (n-1-k)!} \cdot \frac{n-k}{n-k}$$

$$= \frac{(n-1)! k}{k! (n-k)!} + \frac{(n-1)! (n-k)}{k! (n-k)!}$$

$$= \frac{(n-1)! (n-k+k)}{k! (n-k)!} = \frac{(n-1)! n}{k! (n-k)!}$$

$$= \frac{n!}{k! (n-k)!}$$

$$= \binom{n}{k} \quad \text{QED.}$$

Review: how to change the INDEX of summation

Show:  $\sum_{k=1}^n a_{k-1} = \sum_{j=0}^{n-1} a_j$

Technique:  $\sum_{k=1}^n a_{k-1} = \sum_{k=1}^{k=n} a_{k-1}$  put  $j = k-1$   
 $\therefore k = j+1$

$$= \sum_{\substack{j+1=n \\ j+1=1}} a_j$$

$$= \sum_{j=0}^{j=n-1} a_j$$

$$= \sum_{j=0}^{n-1} a_j$$

QED.

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Thm: (Binomial Theorem)

$$(p+q)^n = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k}$$

Prf: (by induction on  $n=1,2,3,\dots$ )

Basis:  $n=1 \quad \therefore (p+q)^1 = p+q$

$$\begin{aligned} \sum_{k=0}^1 \binom{1}{k} p^k q^{1-k} &= \binom{1}{0} p^0 q^1 + \binom{1}{1} p^1 q^0 \\ &= p+q \end{aligned}$$

QED (basis)

Induction Step

Induction hypothesis:  $(p+q)^{n-1} = \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{(n-1)-k}$

$$\therefore (p+q)^n = (p+q)(p+q)^{n-1}$$

$$\begin{aligned} &\stackrel{\text{IH}}{=} (p+q) \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{(n-1)-k} \\ &= \sum_{k=0}^{n-1} \binom{n-1}{k} p^{k+1} q^{n-k-1} + \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{n-k} \end{aligned}$$

$$= \underbrace{\sum_{j=1}^n \binom{n-1}{j-1} p^j q^{n-j}}_{j=k+1 \quad \therefore k=j-1} + \underbrace{\sum_{j=0}^{n-1} \binom{n-1}{j} p^j q^{n-j}}_{j=k}$$

$$= \left[ \underbrace{\binom{n-1}{n-1}}_{=1} p^n q^{n-n} + \sum_{j=1}^{n-1} \binom{n-1}{j-1} p^j q^{n-j} \right] +$$

$$\left[ \sum_{j=1}^{n-1} \binom{n-1}{j} p^j q^{n-j} + \underbrace{\binom{n-1}{0}}_{=1} p^0 q^{n-0} \right]$$

$$= p^n + \sum_{j=1}^{n-1} \binom{n-1}{j-1} p^j q^{n-j} + \sum_{j=1}^{n-1} \binom{n-1}{j} p^j q^{n-j} + q^n$$

$$= p^n + \sum_{j=1}^{n-1} \left[ \underbrace{\binom{n-1}{j-1} + \binom{n-1}{j}}_{=\binom{n}{j}} \right] p^j q^{n-j} + q^n$$

$$\begin{aligned}
&= p^n + \sum_{j=1}^{n-1} \binom{n}{j} p^j q^{n-j} + q^n \\
&= \binom{n}{n} p^n q^{n-n} + \sum_{j=1}^{n-1} \binom{n}{j} p^j q^{n-j} + \binom{n}{0} p^0 q^{n-0} \\
&= \sum_{k=0}^n \binom{n}{k} p^k q^{n-k}
\end{aligned}$$

QED.

### 3 Corollaries

1.  $p = q = 1 \quad \therefore \sum_{k=0}^n \binom{n}{k} = 2^n$

since  $2^n = (1+1)^n = \sum_{k=0}^n \binom{n}{k} 1^k 1^{n-k} = \sum_{k=0}^n \binom{n}{k}$

2.  $p + q = 1 \quad w/ \quad p \geq 0, q \geq 0 \quad (\because q = 1 - p)$

$$\begin{aligned}
\therefore \sum_{k=0}^n P(X=k) &= \sum_{k=0}^n \overset{\text{binomial}}{b(n, k, p)} = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k} \\
&= \overset{\text{BT}}{(p+q)^n} = 1^n = 1 \quad (w/ \quad b(n, k, p) = \binom{n}{k} p^k q^{n-k})
\end{aligned}$$

$\therefore$  Binomial  $b(n, k, p)$  is a probability density

3.  $p = -1$  and  $q = +1$

$$\begin{aligned}
\therefore \sum_{k=0}^n (-1)^k \binom{n}{k} &= 0 \\
&= \binom{n}{0} - \binom{n}{1} + \binom{n}{2} - \binom{n}{3} + \dots + (-1)^n \binom{n}{n}
\end{aligned}$$